

**MASTER'S DEGREE IN PSYCHOLOGY**  
CLINICAL AND HEALTH PSYCHOLOGY

# **Psychometric properties of the European Portuguese Internet Gaming Disorder Test (IGD-20): relation with friendships and physical exercise**

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(IGD-20): relation with friendships and physical exercise**

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Dissertation presented in the Master's Degree in Psychology, Specialization in Clinical and Health Psychology, Faculty of Psychology and Education Sciences of the University of Porto, tutored by Professor Doctor *Fernando Barbosa* (FPCEUP) and co-tutored by Professor Doctor *Artemisa Rocha Dores* (ESS-P.Porto).



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## Resumo

Este estudo visou validar o *Internet Gaming Disorder Test* (IGD-20; Pontes et al., 2014), um instrumento para avaliação da Perturbação de Jogo pela Internet (PJI), em Portugal. Adicionalmente, testaram-se as correlações entre o nível de PJI e número de amigos (NA), frequência de encontros sociais (FES) e exercício físico (EF).

Um total de 556 participantes, com média de idades de 23.3 ( $DP = 6.4$ ), dos quais 61.7% homens, responderam a um protocolo online constituído por questões sociodemográficas, sobre sintomas de ansiedade e depressão, frequência de jogo, NA, FES, horas semanais de EF, PJI e Perturbação de Jogo (PJ). Esta amostra permitiu calcular as consistências internas e realizar Análises Fatoriais Confirmatórias (AFC) com o modelo original e com uma versão curta, o *Internet Gaming Disorder-10 Short Version* (IGD-10SV), de três fatores (Faraci et al., 2023). Utilizou-se uma subamostra de 379 participantes ( $M_{idade} = 23.5$ ,  $DP = 6.4$ ; 61.2% homens) cujo término do protocolo permitiu estimar a validade convergente e de critério, a fidelidade teste-reteste, e testar a relação das pontuações do IGD-20 com as variáveis antes indicadas.

As AFC mostraram um melhor ajuste ao modelo de três fatores, favorecendo uma versão mais curta. O IGD-10SV demonstrou valores moderados de consistência interna, boas validades convergente e de critério e boa fidelidade teste-reteste. Adicionalmente, a pontuação do IGD-10SV demonstrou-se negativamente correlacionada com o NA e com a FES, mas não com o EF.

Apesar das características psicométricas genericamente boas do teste, a ambiguidade da sua estrutura fatorial requer estudo adicional.

*Palavras-chave:* perturbação de jogo pela internet, IGD-20 test, IGD-10SV, propriedades psicométricas, dependência a videojogos

## Abstract

This study aimed to validate the Internet Gaming Disorder Test (IGD-20; Pontes et al., 2014), an instrument for the assessment of Internet Gaming Disorder (IGD), in Portugal. Additionally, the correlations between the level of IGD and number of friends (NF), social gathering frequency (SGF) and physical exercise (PE) were tested.

A total of 556 participants with a mean age of 23.3 ( $SD = 6.4$ ), of which 61.7% were men, answered an online protocol composed of sociodemographic questions, about anxiety and depression symptoms, gaming frequency, NF, SGF, weekly hours of PE, IGD and Gaming Disorder (GD). This sample allowed the calculation of internal consistencies and computation of Confirmatory Factor Analyses (CFA) with the original model and with a short version, the Internet Gaming Disorder-10 Short Version (IGD-10SV), of three factors (Faraci et al., 2023). A subsample of 379 participants ( $M_{age} = 23.5$ ,  $SD = 6.4$ ; 61.2% men) that finished the protocol was used to estimate convergent and criterion validity, test-retest reliability, and test the correlation of IGD-20 scores with the previously indicated variables.

The CFAs showed a better fit for the three-factor model, favoring a shorter version. The IGD-10SV showed moderate values of internal consistency, good convergent and criterion validity, and good test-retest reliability. Additionally, IGD-10SV scores showed to be negatively correlated with NF and SGF, but not with PE.

Despite the scale's good overall psychometric properties, the ambiguity of its factorial structure warrants further study.

*Keywords:* internet gaming disorder, IGD-20 test, IGD-10SV, psychometric properties, videogame addiction

## Résumé

Cette étude a pour objectif valider le *Internet Gaming Disorder Test* (IGD-20 ; Pontes et al., 2014), un instrument d'évaluation du trouble du jeu sur Internet (IPD), au Portugal. En outre, les corrélations entre le niveau de PJI et le nombre d'amis (NA), la fréquence des rencontres sociales (FES) et l'exercice physique (EF) ont été testées.

Au total, 556 participants, âgés en moyenne de 23,3 ans (écart-type = 6,4), dont 61,7 % d'hommes, ont répondu à un protocole *online* comprenant des questions sociodémographiques, des symptômes d'anxiété et de dépression, la fréquence des jeux, NA, FES, le nombre d'heures par semaine d'EP, PJI et le trouble du jeu (PJ).

Cet échantillon a permis de calculer les cohérences internes et de réaliser des analyses factorielles confirmatoires (AFC) avec le modèle original et avec une version courte, l'*Internet Gaming Disorder-10 Short Version* (IGD-10SV), à trois facteurs (Faraci et al., 2023). Un sous-échantillon de 379 participants ( $M_{\text{âge}} = 23,5$ , écart-type = 6,4 ; 61,2 % d'hommes) ayant complété le protocole a été utilisé pour estimer la validité convergente et la validité de critère, la fiabilité *test-retest* et pour tester la relation entre les scores de l'IGD et les variables susmentionnées.

Les AFC ont montré une meilleure adéquation au modèle à trois facteurs, favorisant une version plus courte. L'IGD-10SV a montré des valeurs de cohérence interne modérées, une bonne validité convergente et critérielle et une bonne fiabilité *test-retest*. En outre, le score IGD-10SV était négativement corrélé à la NA et à la FES, mais pas à l'EF.

Malgré les bonnes caractéristiques psychométriques générales du test, l'ambiguïté de sa structure factorielle nécessite une étude plus approfondie.

*Mots-clés:* trouble du jeu sur internet, IGD-20 test, IGD-10SV, propriétés psychométriques, dépendance aux jeux vidéo

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## Introduction

Gaming is one of the most prevalent leisure activities in the world, involving approximately 3.2 billion active gamers in 2022 (Warman et al., 2022). Notwithstanding their potential benefits to cognitive performance (Boot et al., 2008; Green & Bavelier, 2015; Kovbasiuk et al., 2022) and therapeutic application (Bocci et al., 2023; Colder Carras et al., 2018; Mahon-Daly & Jeyasingam, 2021), video games, when played in excess, can contribute among vulnerable individuals to the development of pathological habits that may have detrimental effects on mental and physical health (Ko, 2014; Ko et al., 2020).

Researchers became more interested in the topic after the two current nosological manuals proposed criteria for labelling such problematic gaming behaviors. The 5<sup>th</sup> version of the Diagnostic and Statistical Manual of Mental Disorders (DSM-5; American Psychiatric Association [APA], 2013) introduced *Internet Gaming Disorder* (IGD) as a condition for further study, defining it as the “persistent and recurrent use of the Internet to engage in games, often with other players, leading to clinically significant impairment or distress” (pp. 795). According to APA (2013):

At least five of the following criteria need to be present for 12 months:

(1) preoccupation with Internet games, in anticipation of, or after, playing, (2) withdrawal symptoms such as irritability, anxiety or sadness when Internet gaming is taken away, (3) tolerance or the need to spend increasing amounts of time engaged in Internet Games, (4) unsuccessful attempts to control the participation in Internet games, (5) loss of interests in previous hobbies and entertainment as a result of, and with the exception of, Internet games, (6) continued excessive use of Internet games despite knowledge of psychosocial problems, (7) has deceived family members, therapists, or others regarding the amount of Internet gaming, (8) use of Internet games to escape or relieve a negative mood such as helplessness, guilt and anxiety, and (9) has jeopardized or lost a significant relationship, job, or educational or career opportunity because of participation in Internet games. (pp. 795-796)

Similarly, the 11<sup>th</sup> revision of the International Classification of Diseases (ICD-11; World Health Organization [WHO], 2022) added *Gaming Disorder* (GD) under the category of *disorders due to addictive behaviors* and defined it as a “pattern of persistent

or recurrent gaming behaviors, which may be *online* (i.e., over the internet) or *offline*". Individuals who exhibit such a pattern for at least 12 months may experience difficulties controlling their gaming behavior, which ends up taking precedence over other areas of life, leading to interpersonal and relational problems.

The estimated global prevalence of gaming addiction may lie between 1.4% and 3.3% (H. S. Kim et al., 2022) or 2.0% and 3.1% (Stevens et al., 2021), and individuals with a diagnosis are mostly men (Darvesh et al., 2020; Mihara & Higuchi, 2017). However, researchers do not consider these estimates very conclusive, highlighting limitations in the studies, such as the use of different conceptualizations, discrepancies in assessment tools, and low sample representativeness (H. S. Kim et al., 2022; Kircaburun et al., 2020; Stevens et al., 2021).

The use of different conceptualizations derives from an extensive debate in the scientific community (Griffiths et al., 2016; Király et al., 2015; Petry et al., 2014, 2016) that spans the legitimacy of the disorder (Starcevic, 2013; Van Rooij & Kardefelt-Winther, 2017), the pertinence of proposed criteria (Kuss et al., 2017) and the addictive potential of offline games (Krossbakken et al., 2017; Montag et al., 2021). In attempting to reach a consensus, researchers have appealed to research based in robust methodologies, theories and validated diagnostic tools (Griffiths et al., 2014, 2016; Kuss et al., 2017; Petry et al., 2014, 2016; Van Rooij & Kardefelt-Winther, 2017).

Notwithstanding the previously mentioned conceptual and methodological issues, there is also a problem related with a high comorbidity among individuals diagnosed with IGD, specifically depression (Laconi et al., 2017; Ostinelli et al., 2021), anxiety (Colder Carras et al., 2020; Huang et al., 2022), attention deficit/hyperactivity disorder (Barrangou-Pouey-Darlas et al., 2022; Dullur et al., 2021), social phobia (Gioia et al., 2022) and obsessive-compulsive disorder (González-Bueso et al., 2018). Considering that there is a lack of clarity regarding the directionality between the comorbidities and IGD (Colder Carras et al., 2020; Dullur et al., 2021; González-Bueso et al., 2018; Kircaburun et al., 2020; Laconi et al., 2017), it becomes difficult to determine whether IGD is a primary disorder with comorbidities, or the consequence of pre-existing pathologies.

Anxiety and depression appeared to be highly associated with IGD in a study by González-Bueso et al. (2018). Additionally, other studies have found positive and predictive relationships between the symptoms of both disorders and IGD (Gentile et al., 2011; S. Kim et al., 2022; Teng et al., 2021), as well as a mediation effect of IGD on anxiety and depression levels (Stavropoulos et al., 2022).

Concomitantly, IGD is associated with risk factors and psychosocial and physical consequences. However, not only are the relationships between variables unclear, but there is also a lack of distinction over what constitutes risk and consequence of IGD (Mestre-Bach et al., 2022; Paulus et al., 2018). The authors propose that the blurred lines between risk factors and consequences of the disorder are due to the lack of longitudinal studies, small sample sizes and an exacerbating cycle of the variables that first originated IGD.

A systematic review by Paulus et al. (2018) proposed a comprehensive model of etiological factors related to IGD. While their review covers topics beyond the scope of this study, it is important to highlight the associations found between IGD and the following variables: stress; affect dysregulation, or the tendency for negative, depressive and anxious affective states; low self-esteem; solitude; social isolation; reduced interpersonal relationships and lack of real-world friends; social inhibition; compromised social functioning; low quality relationships with parents and peers; and low involvement in leisure activities. Similarly, a recent meta-analysis identified the following risk factors: stress, average gaming time, family dysfunction, weak academic performance, engaging in bullying, either as a victim or aggressor, interpersonal problems, hyperactivity/attention deficit, anxiety, depression, emotional distress, and low self-esteem (Gao et al., 2022). On the other hand, a longitudinal study found a negative association between IGD and psychosocial well-being, but not the opposite, suggesting that the disorder may be a maladaptive response that leads to low self-esteem, limited social support and dissatisfaction with life (Teng et al., 2020).

Researchers hypothesize that the varying results across studies regarding the quality of interpersonal relationships among pathological gamers may be attributed to the replacement of real-world bonds with virtual-world ones, which may be safer and less anxiety-inducing (Király et al., 2023; Sublette & Mullan, 2012). In other words, while some gamers remain isolated and devoid of real-world friendships, other individuals find gratification in online bonds, reporting less dissatisfaction compared to the first group. Moreover, virtual-world relationships may contribute to addiction, as gamers play more in order to fulfil social needs that are unmet in real life (Heng et al., 2021; Király et al., 2023; Szolin et al., 2022).

Excessive gaming can also lead to negative physical consequences such as: visual problems, musculoskeletal problems, auditory hallucinations, enuresis, encopresis, sore tendons, tenosynovitis, skin blisters, calluses, hand-arm vibration syndrome, peripheral neuropathy, epileptic seizures, overweight, and obesity (Ayenigbara, 2018; Paulus et al.,

2018). Obesity and overweight may be associated with this disorder due to a sedentary lifestyle, typical among individuals who spend a lot of time sitting in front of a screen (Bullock et al., 2017). Consequently, a negative association between physical exercise (PE) and the severity of IGD is expected. However, the results from studies on this topic have been inconclusive (Huard Pelletier et al., 2020).

Three studies found no association between PE, body mass index (Bickham et al., 2013; Wack & Tantleff-Dunn, 2009), and IGD (Männikkö et al., 2015). On the other hand, both negative associations (Ballard et al., 2009) and reciprocal causality (Henchoz et al., 2016) were found between PE and IGD. Finally, Puolitaival et al. (2020) reported a higher incidence of sedentary behaviors, including lower practice of PE, among individuals who spent more time playing videogames.

As stated previously, researchers are attempting to gather data on the prevalence of IGD, to establish a consensus on its diagnostic criteria, and to develop reliable assessment tools (King et al., 2020; Petry et al., 2014). Consequently, several tools that attempt to assess IGD and standardize its screening methods have emerged. According to a qualitative content validity analysis by Karhulahti et al. (2021), there were 17 instruments designed for assessing IGD based on DSM-5 (APA, 2013) or ICD-11 (WHO, 2022) criteria. King et al. (2020) reported a total of 32 different instruments for assessing GD or IGD, regardless of their theoretical background. For the sake of comparing psychometric properties, the 20 instruments from these studies that were based on criteria from the two nosological manuals were selected and presented in Table 1.

**Table 1**  
*IGD and GD assessment instruments and psychometric properties*

<b>Instrument (authors)</b>	<b>IN</b>	<b>IC (<math>\alpha</math>)</b>	<b>CV (<math>r</math>)</b>	<b>CRV (<math>r</math>)</b>	<b>TRR (<math>r</math>)</b>
GAIA (Wong & Hodgins, 2014)	26	.94	-	SSBA: .79, YIAT: .80	-
IGDC (Petry et al., 2014)	9	.69 – .86	.24 – .59	IGCS: .51, GAS- 7: .76	-
IGD-20 (Pontes et al., 2014)	20	.85 – .92	.27 – .77	BAM-VG: .44, DSM-5 criteria & IGDS9-SF: .82	-
IGDS (Lemmens et al., 2015)	9/ 27	.67 – .95/.94 – .95 <sup>a</sup>	.22 – .66/.23	YDQ: .39, PIGDS: .78/ - <sup>a</sup>	-

Instrument (authors)	IN	IC ( $\alpha$ )	CV ( $r$ )	CRV ( $r$ )	TRR ( $r$ )
	a		-.56 <sup>a</sup>		
IGDS9-SF (Pontes & Griffiths, 2015)	9	.81 – .86	.32 – .55	GAS-7: .40, YIAT: .83	.87 (2 weeks)
GAIT (Vadlin et al., 2015)	15	.89 – .91	.48	GAS-7: .83	-
CSAS (Rehbein et al., 2015)	18	.92 – .94	-	-	-
PIE-9 (Pearcy et al., 2016)	9	.86 – .89	-	PVP: .43, IGD-20: .64	.77 (2 weeks)
BAM-VG (Sanders & Williams, 2016)	19	.87	-	IGD-20: .44	.73 (90 days)
IGDT-10 (Király et al., 2017)	10	.62 – .87	.03 (non-sig) – .45	PIUQ-6: .53, POGQ: .77	-
VASC (Yılmaz et al., 2017)	21	.89	-	-	-
CVAT 2.0 (Van Rooij et al., 2017)	11	-	-	-	-
C-IGDS (Sigerson et al., 2017)	9	.91	.41	CES-D: .62, SIAS: .37	-
IGUESS (Jo et al., 2018)	9	.85 – .94	IGD > non-IGD	YIAT: .90	-
IGD Scale (Finslerås et al., 2019)	9	.78	-	GASA: .71	-
IGDS-23 (Borges et al., 2019)	23	.82	-	-	-
GADIS-A (Paschke et al., 2020)	10	.91	.42 <sup>b</sup>	IGDS: .7, PIGDS: .54	-
GDHGS (Balhara et al., 2020)	4	.91	-	IGDS9-SF: .88	-
TIGTOC (Jo et al., 2020)	3	.81	-	YIAT: .66, CDI: .41, K-ARS: .003	-
GDT (Pontes et al., 2021)	4	.84	- <sup>c</sup>	IGDS9-SF: .83	-

*Note.* Convergent validity was measured via bivariate correlation with gaming hours per week, when available. A dash was used whenever data was not reported. IN = Item Number; IC = Internal Consistency; CV = Convergent Validity; CRV = Criterion Validity; TRR = Test-Retest Reliability.

List of tools: BAM-VG: Behavioral Addiction Measure – Video Gaming; CDI: Child Depression Inventory (Kovacs, 2015); CES-D: Center for Epidemiologic Studies Depression Scale-Short Form (J. C. Cole et al., 2004); C-IGDS: Chinese Internet Gaming Disorder Scale; CSAS: Video Game Dependency Scale; CVAT 2.0: Clinical Video game Addiction Test 2.0; GAIA: Game Addiction Inventory for Adults; GADIS-A: Gaming Disorder Scale for Adolescents; GAIT: Gaming Addiction Identification Test; GASA: Gaming Addiction Scale for Adolescents (Lemmens et al., 2009); GAS-7: Game Addiction Scale (Lemmens et al., 2009); GDHGS: Gaming Disorder and Hazardous Gaming Scale; IGDC: Internet Gaming Disorder Checklist; GDT: Gaming Disorder Test; IGCS: Internet Gaming Cognition Scale (King & Delfabbro, 2016); IGD-20: Internet Gaming Disorder-20 Test; IGDS: Internet Gaming Disorder Scale; IGDS9-SF: Internet Gaming Disorder Scale-Short-Form; IGDT-10: Ten-Item Internet Gaming Disorder Test; IGUIESS: Internet Game Use-Elicited Symptom Screen; K-ARS: Korean ADHD Rating Scale (DuPaul, 1991); PIE-9: Personal

Internet Gaming Disorder Evaluation 9; PIGDS: Parental version of the Internet Gaming Disorder Scale (Wartberg et al., 2019); PIUQ-6: Problematic Internet Use Questionnaire-6 items (Demetrovics et al., 2016); POGQ: Problematic Online Gaming Questionnaire (Demetrovics et al., 2012); PVP: Problematic Video game Playing Scale (Tejeiro Salguero & Morán, 2002); SIAS: Social Interaction Anxiety Scale-Short form (Peters et al., 2012); SSBA: Screener for Substance and Behavioral Addictions (Schluter et al., 2018); TIGTOC: The Three-Item Gaming disorder Test–Online Centered; VASC: Video Game Addiction Scale for Children; YDQ: Young Diagnostic Questionnaire (Young, 1998); YIAT: Young Internet Addiction Test (Young, 1998).

With the exception of the following tools: C-IGDS, IGD Scale, IGDS-23, GDT, GADIS-A, GDHGS and TIGTOC whose psychometric qualities were retrieved from their respective development studies, the table and its values were adapted from "Screening and assessment tools for gaming disorder: A comprehensive systematic review." by King, D. L., Chamberlain, S. R., Carragher, N., Billieux, J., Stein, D. J., Mueller, K., Potenza, M. N., Rumpf, H., Saunders, J. B., Starcevic, V., Demetrovics, Z., Brand, M., Lee, H. Y., Spada, M. M., Lindenbergh, K., Wu, A. M. S., Leménager, T., Pallesen, S., Achab, S., . . . Delfabbro, P., 2020, *Clinical Psychology Review*, 77, 101831 (<https://doi.org/10.1016/j.cpr.2020.101831>). Copyright by Creative Commons Attribution Non-Commercial-No Derivatives License (CC BY NC ND).

<sup>a</sup> Refers to two versions of the scale with 9 and 27 items, respectively. <sup>b</sup> Gaming hours per day. <sup>c</sup> Data that demonstrates convergent validity: AVE  $\geq$  .50; Composite Reliability  $\geq$  .70; No cross-loadings. For more information refer to Pontes et al. (2021).

The *Internet Gaming Disorder Test* (IGD-20 Test; Pontes et al., 2014) was the first screening tool designed for the assessment of IGD, based on DSM-5 (APA, 2013) criteria. The original version showed good internal consistency and concurrent validity, exhibiting correlations that ranged from weak to strong with analogous tools (BAM-VG = .44; IGDS9-SF = .82; King et al., 2020). Regarding convergent validity, associations with weekly gaming hours ranged from weak ( $r = .27$ ) to strong ( $r = .77$ ). B.-N. Kim (2019) suggests that the variations observed in convergent validity may be related with the choice of the variable *weekly gaming time*, which is highly dependent on sample characteristics.

The *IGD-20 Test* was based on a six-factor model for addiction proposed by Griffiths (2005) that is theoretically equivalent to the nine DSM-5 criteria (Pontes et al., 2014). The six factors that compose the model are as follows: *salience*, *mood modification*, *tolerance*, *withdrawal*, *conflict* and *relapse*. Salience is related to the intruding importance that the addictive activity represents in daily life, oftentimes dominating thoughts, feelings and behaviors. Mood modification concerns the numbing or pleasant feelings and/or emotional states that are elicited while engaging in the addictive behavior. Tolerance is the ever growing need to increase the amounts of a particular addictive activity in order to experience previously felt pleasant or numbing effects. Withdrawal refers to “unpleasant feeling states and/or physical effects that occur when the activity” is halted or reduced (Griffiths, 2005, p.194). Conflict relates to any altercations and discordance that arise within the individual and between the individual and those around them, as a result of their addictive behavior. Finally, relapse pertains to the tendency to quickly resume patterns of

addiction even after a period of abstinence or control.

While in the scale's development study a Confirmatory Factor Analysis (CFA) demonstrated a good fit for Griffith's addiction model (Pontes et al., 2014), its subsequent validations reported a variety of results. Specifically, other acceptable six-factor structures were found (Fuster et al., 2016; B.-N. Kim, 2019; Pontes et al., 2014) as well as very good five-factor (Cakiroglu & Soylu, 2019; Grajewski & Dragan, 2021; Yu et al., 2019) and one-factor (L. I. Andrade et al., 2022; Hawi & Samaha, 2017) structures. In both unifactorial solutions, an Exploratory Factor Analysis (EFA) was utilized alongside the CFA, which may explain the results. Additionally, all five-factor structures opted to combine the factors salience and tolerance, as they were highly correlated (Cakiroglu & Soylu, 2019; Grajewski & Dragan, 2021; Yu et al., 2019).

The recent Italian validation concluded that the factorial structure of the scale was fundamentally flawed to some extent, leading to very high correlations among latent variables and mixed results of factorial internal consistency (Faraci et al., 2023). As a result of their study, the scale was cut down to 10 items and its factorial structure was changed to three factors: (a) *prevalence*, or the negative contextual impact that the time devoted to pathological gaming has on individuals; (b) *withdrawal*, defined like Griffiths (2005); and (c) *craving*, which is based on Caretti et al.'s (2008) work, is described as the "longing desire" to play videogames that pervades day-to-day life. The new scale, or *Internet Gaming Disorder Scale-10 Short Version (IGD-10SV)*, remained faithful to the theoretical framework of addiction while showing an overall good fit and solving the high intercorrelations among factors present in all of the previous validations.

Finally, the IGD-20 has been translated and validated in Spain (Fuster et al., 2016), for the Arabic population (Hawi & Samaha, 2017), Korea (B.-N. Kim, 2019), China (Yu et al., 2019), Turkey (Cakiroglu & Soylu, 2019), Poland (Grajewski & Dragan, 2021), Ecuador (L. I. Andrade et al., 2022) and Italy (Faraci et al., 2023).

In Portugal, to date, only two gaming addiction scales have been validated for the European Portuguese population (Beirão et al., 2019). These scales are the Videogame Addiction Scale (Escala de Adição ao Jogo Electrónico [EAJE]; Barroso et al., 2011), and the IGDS9-SF (Pontes & Griffiths, 2016). Additionally, both scales were validated in an adolescent sample, which only represents a portion of the population affected by IGD (Gao et al., 2022). Furthermore, the EAJE does not reflect the DSM-5 criteria (APA, 2013) as it was developed before their release. Lastly, the IGD-20, in theory, may reduce measurement error (Böckenholt & Lehmann, 2015) as it consists of 20 items, whereas the

IGDS9-SF only comprises nine. Additionally, its factorial structure could provide further insights to the IGD debate.

Given the extant literature, this study aimed to validate the IGD-20 for the European Portuguese population and examine its psychometric properties. Such validation intends to tackle the lack of IGD assessment tools in Portugal and provide a new screening tool for researchers and mental health professionals. Simultaneously, this work intended to estimate the correlations among the variables under study, establishing the following hypotheses:

- a) The number of friends (NF) will be negatively correlated with IGD scores;
- b) The social gathering frequency (SGF) will be negatively correlated with IGD scores;
- c) The weekly hours of PE will be negatively correlated with IGD scores.

## **Methods**

### **1. Sample**

Participants were recruited through different methods, as detailed in the procedures section below.

Only participants who were (a) at least 18 years of age, (b) of European Portuguese nationality and who (c) played on electronic platforms (d) for more than 0 days per week were included in the analyses.

According to Kyriazos (2018), in order to conduct a successful CFA, the ratio of cases (N) to parameters (q) should be at least 10:1. Since the IGD-20 questionnaire consists of 20 questions, a minimum of 200 participants was considered necessary for factor analysis. For test-retest reliability, a minimum sample of 100 respondents was considered to achieve the desired results (Kennedy, 2022).

All collected responses were screened for missing values. Any participants who did not fully complete the IGD-20 were excluded, leaving 633 participants in the sample. Following this step, every participant who did not meet inclusion criteria was filtered out of all analyses, leaving a final sample of 556 respondents, aged from 18 to 55 years old

(see full characteristics in Table 2).

Additionally, a subsample of 379 participants was established after filtering 177 respondents who did not complete the full data collection protocol (i.e., those who only completed the IGD-20). Participants of this subsample ranged in age from 18 to 55 years old (see Table 2).

Table 2 contains the descriptive statistics regarding variables of interest that characterize the sample and subsample (means and standard deviations or frequencies, depending on the variables). For more information on the sample's national representativeness, Appendix A can be consulted.

**Table 2**

*Descriptive statistics of the total sample, subsample used in the CFA, and subsample that completed the retest*

Variable	Main Sample (N = 556)		Subsample (n = 379)		Retest Subsample (n = 112)	
	<i>M</i>	<i>SD</i>	<i>M</i>	<i>SD</i>	<i>M</i>	<i>SD</i>
Age	23.3	6.36	23.5	6.39	23.1	5.72
N. years of formal education	14.2	2.14	14.3	2.26	14.5	2.44
Daily gaming hours	2.28	1.67	2.28	1.71	2.46	1.95
Weekly gaming days	4.22	2.03	4.29	2.04	4.62	2.02
Physical exercise	2.97	3.63	3.16	3.87	2.94	3.75
	<i>n</i>	<i>%</i>	<i>n</i>	<i>%</i>	<i>n</i>	<i>%</i>
Gender						
Men	343	61.7	232	61.2	71	63.4
Women	213	38.3	147	38.8	41	36.6
Region of Residence						
Northern	320	57.6	220	58.0	72	64.3
Central	86	15.5	57	15.0	9	8.0
Lisbon and Tagus Valley	62	11.2	44	11.6	14	12.5
Alentejo	33	5.9	20	5.3	6	5.4
Algarve	38	6.8	25	7.0	7	6.3
Madeira	11	2.0	9	2.4	2	1.8
Azores	6	1.1	4	1.1	2	1.8

## 2. Materials

The original data collection protocol was part of a larger study being conducted online by the research team of the Laboratory of Neuropsychophysiology, of the Faculty of Psychology and Education Sciences, University of Porto (FPCEUP), in cooperation with the School of Health, Polytechnic of Porto (ESS-P.PORTO, see Appendix B). Online data collection provides easier access for participants, facilitated advertising and recruitment, automation of certain tasks and shorter time of data collection (Wood & Griffiths, 2007; Wright, 2006). This method is also suitable for studying behavioral addictions, since most gamers are probably familiar with browsing the internet (Wood & Griffiths, 2007).

A select number of materials and measures from the original protocol were used in this study, as described below.

### ***Questionnaire on Sociodemographic Data, Gaming Frequency, Number of Friends (NF), Social Gathering Frequency (SGF), and Physical Exercise (PE)***

The research team developed and administered a questionnaire intended to collect:

- sociodemographic data, including age, gender, district of residence, current degree, current year of study, and number of years of formal education;
- *weekly gaming days* with response options ranging from 0 to 7;
- *daily gaming hours*, answered in an open-ended format limited to numbers only;
- FQ was assessed on a scale from 1 (*none*) to 4 (*many*);
- SGF was answered on a scale from 1 (*none*) to 5 (*very frequently*);
- PE through open text entry.

### ***European Portuguese Version of the Internet Gaming Disorder Test (IGD-20; Pontes et al., 2014)***

The IGD-20 (Pontes et al., 2014) is a self-report tool designed for the assessment of recurrent and persistent IGD, for both online and offline games, over the course of the past 12 months. It totals 20 items, answered on a 5-point Likert scale: 1 (*totally disagree*), 2 (*disagree*), 3 (*neither agree or disagree*), 4 (*agree*) and 5 (*totally agree*). Scores range from 20 to 100, where a higher level indicates higher IGD. In the original version a cut-off of 71 distinguished non-disordered gamers from disordered ones. Moreover, the items are organized by six factors: salience (items 1, 7 and 13), mood modification (items 2, 8 and

14), tolerance (items 3, 9 and 15), withdrawal symptoms (items 4, 10 and 16), conflict (items 5, 11, 17, 19 and 20) and relapse (items 6, 12 and 18). Only items 2 and 19 are inverted.

Additionally, the internal consistency of the original version as revealed by Cronbach's Alpha was .88.

Finally, the original author, who is bilingual, had already directly translated the scale to European Portuguese.

### ***European Portuguese Version of the Gaming Disorder Test (GDT; Pontes et al., 2021)***

The GDT is a short scale that measures the severity of GD, in the past 12 months, based on the ICD-11 criteria (WHO, 2022). It comprises four items, rated on a 5-point scale: 1 (*never*), 2 (*rarely*), 3 (*sometimes*), 4 (*often*) and 5 (*very often*). Scores can range from 4 to 20 with higher scores indicating higher degrees of GD. In order to differentiate disordered gamers from non-disordered gamers, all four items need to be answered with 4 (*often*) and 5 (*very often*).

The GDT was also directly translated by the original author, who is bilingual, and it showed good internal consistency in this sample ( $\omega = .85$ ). The internal consistency coefficient found in this sample is similar to that of the original study, where the Cronbach's Alpha of the overall sample was .84 (Pontes et al., 2021).

### ***Hospital Anxiety and Depression Scale (HADS; European Portuguese validation by Pais-Ribeiro et al., 2007)***

The HADS is a 14 item self-report tool that evaluates depressive and anxious affective states (Zigmond & Snaith, 1983). Its items are divided equally by two subscales, anxiety and depression, and are rated from 0 to 3 on a 4-point scale with a variety of possible answers unique to each item. Odd numbered items belong to the anxiety subscale whereas even numbered items pertain to the depression subscale. The total scores of both subscales can range from 0 to 21 with a score of 11 potentially indicating the presence of a mood disorder. Items 1, 3, 5, 6, 8, 10, 11 and 13 are inverted.

Finally, internal consistency for the anxiety and depression subscales showed good ( $\omega = .83$ ) and acceptable scores ( $\omega = .74$ ), respectively. In the original version, Cronbach's alpha for the anxiety scale was acceptable ( $\alpha = .76$ ) whereas, for the depression scale, it was good ( $\alpha = .81$ ), differing slightly from this sample.

### 3. Data Collection and Analysis Procedures

The surveys were created on Qualtrics, version 04-2023 (Qualtrics, Provo, Utah, USA), and hosted in the same platform from June 2022, for a total of 10 months and 10 days.

A total of 1021 participants were recruited through a link shared on Facebook, Instagram and Portuguese internet gaming forums, communities, and Discord servers. Besides these recruitment methods, Higher Education Institutions in the country were contacted via e-mail and asked to share the survey with their respective academic communities. Finally, a snowball sampling method, in which participants were encouraged to recruit other participants, was employed.

The study was approved by the Ethics Committee of the ESS-P.PORTO (CE0100C, see Appendix C), and followed the European Code of Conduct for Research Integrity (ALLEA, 2023) and the General Data Protection Regulation recently approved for EU countries (Regulation 2016/679). All participants actively accepted the informed consent prior to responding. Additionally, in order to enable the sending of a link for the retest and to pair of both sets of responses, the participants could choose to provide their e-mail addresses. All addresses were then associated with a unique Respondent ID and kept separate from other sensitive data. Subsequently, Microsoft Office Excel was used to execute a command line that mass replaced all e-mails with their corresponding IDs in the retest database. Therefore, the investigators only had to access the association key when a participant indicated a different or misspelled e-mail address.

Eight days after completion of the first survey, a link to the retest of IGD-20 was sent to the e-mail addresses previously provided by the participants. The association file containing all e-mails and IDs was destroyed afterwards, alongside the databases downloaded from Qualtrics.

Regarding statistical analysis, every variable of interest was screened for out-of-range values and outliers. The main sample contained two out-of-range values of 0 in the variable *Age* that were replaced with their most likely values, considering the respondent's number of years of formal education, current degree and current year of study. Moreover, five outliers were discovered in the IGD-20 total scores. In the subsample, the variables IGD-20, daily gaming hours, GDT and HADS depression scale showed three, eight, 10 and four outliers, respectively. The variable PE showed nine outliers which belonged to participants with much longer time devoted to physical exercise (>13 weekly hours) than

the remaining sample. All outliers were considered plausible, expected, and naturally occurring. Additionally, the analyses were run with and without the outliers, but no significant differences were found (see Appendix D).

Not all variables were normally distributed, but none deviated heavily from normality (H.-Y. Kim, 2013, see Table 3).

**Table 3**  
*Skewness and Kurtosis*

Variable	Main Sample ( <i>N</i> = 556)		Subsample ( <i>n</i> = 379)		Retest Sample ( <i>n</i> = 112)	
	Skewness	Kurtosis	Skewness	Kurtosis	Skewness	Kurtosis
IGD-20	0.71	0.28	0.67	-0.00	0.48	-0.80
Daily gaming hours			1.55	3.83		
Weekly gaming days			0.03	-1.37		
GDT			1.43	2.40		
HADS_A			0.46	-0.45		
HADS_D			0.86	0.92		
Physical exercise			1.60	3.29		

*Note.* Blank spaces = not applicable in the analyses. HADS\_A = anxiety subscale of HADS; HADS\_D = depression subscale of HADS; IGD-20\_II = Retest phase of the IGD-20.

To test the original six-factor model of the IGD-20, a CFA using Maximum Likelihood with Robust Standard Errors (MLR) was performed. Following model rejection, the modification indices (MI) were analysed and an attempt at respecifying the model was made. Given the rejection of the respecification (i.e., redefining the relations between the variables contained in the model; Kline, 2023), the three-factor solution proposed recently by Faraci et al. (2023) was tested using the Unweighted Least Squares (ULS) estimation method. The ULS was chosen over the Diagonal Weighted Least Squares (DWLS) estimation method because it showed better performance for ordinal measures (Forero et al., 2009). All global fit indices were judged based on commonly used thresholds: Comparative Fit Index (CFI) and Tucker Lewis Index (TLI) were considered acceptable within the .90-.95 margin and excellent when >.95. Root Mean Square Error of Approximation (RMSEA) and Standardized Root Mean Residual (SRMR) were considered acceptable when values lay within the .06-.08 range, whereas values < .05 were considered excellent (Hu & Bentler, 1999). The assessment of overall and local fit of the models

followed the guidelines proposed by Kline (2023).

To investigate the internal consistency of the scale, a reliability analysis using McDonald's Omega and Cronbach's Alpha was conducted on both models. Cronbach's Alpha is more commonly used, but McDonald's Omega is preferred (Hayes & Coutts, 2020; McNeish, 2018).

In order to verify convergent and concurrent validity, as well as the test-retest reliability, correlational analyses were conducted using Pearson's  $r$ . Convergent validity was measured by correlating the total scores of the IGD-10SV with both HADS subscales and the total scores of the GDT. Concurrent validity was measured by correlating the total score of the IGD-10SV with both gaming frequency variables, namely daily gaming hours and weekly gaming days. Since all variables contained outliers, Spearman's Rho was also estimated. According to Khamis (2008), Spearman's coefficient is more adequate when one or more assumptions are violated or outliers are present, even though Puth et al. (2015) have argued that the "robustness of approaches based on Pearson's measure makes this approach unnecessary" (p. 78). The results from both coefficients showed very similar results (see Appendix D).

Finally, Kendall's Tau-b was used to test the first two hypotheses. In concordance with Woods (2007), the Kendall's Tau-b is more adequate than Spearman's Rho when conducting a correlation between a continuous variable and a categorical variable with four or five levels. For the third hypothesis, Pearson's  $r$  was used instead, since both of the involved variables were quantitative.

Both RStudio for R version 4.3.0 (R Core Team, 2023) with the lavaan package (version 0.6-16; Rosseel, 2012) and JASP (version 0.17.3; JASP Team, 2023) were used for the CFA's. Descriptive analyses, internal consistency, convergent validity, criterion validity, test-retest reliability and hypotheses testing were all performed in Jamovi (version 2.3.21; The jamovi project, 2023).

## Results

### 1. Confirmatory Factor Analysis

The six-factor model did not produce an acceptable fit,  $\chi^2_{(155, N = 556)} = 498.07, p <$

.001; CFI = .899; TLI = .876; RMSEA = .068 (90% CI: .062 — .075); SRMR = .061. An analysis of the MI led to the error correlation of items 6 and 18 (MI = 32.10), because they belonged to the same latent factor and could correlate due to similar wording.

The correlation of the errors of these items produced a marginally acceptable fit:  $\chi^2_{(154, N = 556)} = 466.62, p < .001$ ; CFI = .908; TLI = .886; RMSEA = .066 (90% CI: .059 — .072); SRMR = .054. However, this model was rejected due to the high chi-square value (although a  $\chi^2/df \leq 3$  indicates an acceptable fit according to Kline, 1998) in combination with three inadmissible correlations of  $> 1$  among factors (see Table 4), which produced a covariance matrix not positive-definite.

**Table 4**  
*Pattern Coefficients and estimated latent factor correlation matrix*

Items	Latent Factors					
	Saliency	Mood	Tolerance	Withdrawal	Conflict	Relapse
	<b>Modification</b>					
Item 1	.51					
Item 7	.65					
Item 13	.60					
Item 2		.41				
Item 8		.87				
Item 14		.84				
Item 3			.38			
Item 9			.72			
Item 15			.53			
Item 4				.70		
Item 10				.66		
Item 16				.75		
Item 5					.62	
Item 11					.62	
Item 17					.59	
Item 19					.43	
Item 20					.71	
Item 6						.61
Item 12						.48
Item 18						.61
Saliency	-					
Mood Modification	.66	-				
Tolerance	<b>1.03</b>	.68	-			

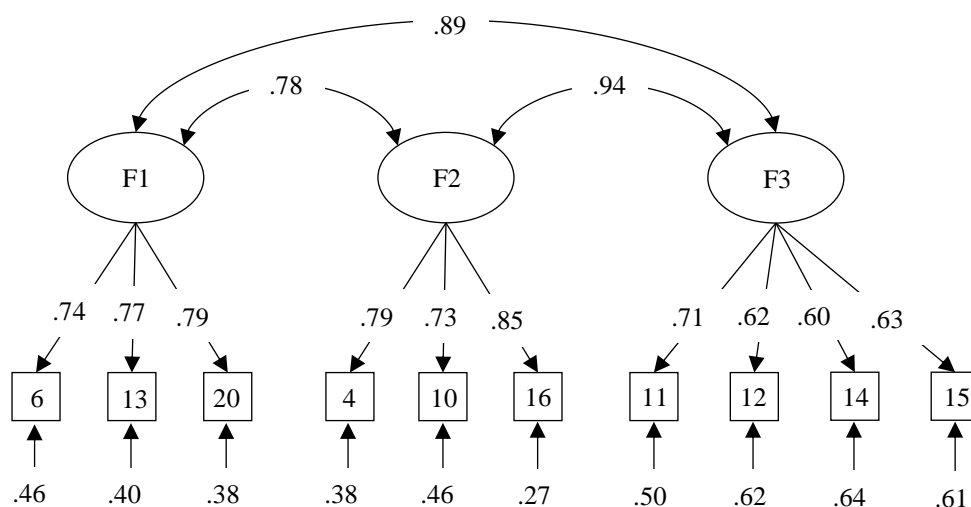
Items	Latent Factors					
	Salience	Mood	Tolerance	Withdrawal	Conflict	Relapse
	<b>Modification</b>					
Withdrawal	.90	.57	.93	-		
Conflict	.86	.48	.77	.80	-	
Relapse	<b>1.04</b>	.62	.90	.87	<b>1.06</b>	-

Similar problems were highlighted by Faraci et al. (2023) who found that a three-factor solution best fit their data. Ultimately, a CFA using an ULS estimation method was performed with that model, which yielded acceptable results on a global level:  $\chi^2_{(32, N = 556)} = 93.57, p < .001$ ; CFI = .989; TLI = .985; RMSEA = .059 (90% CI: .045 — .073); SRMR = .055. Figure 1 shows the accepted three-factor model with its pattern coefficients. Except for the high correlation between withdrawal and craving, the model was considered satisfactory. Furthermore, an analysis of the residuals' matrix showed six values exceeding .10 (see Table 5), meaning that the model does not account for all of the observed correlations. However, because these items derive from the same instrument, the results are somewhat expected (D. A. Cole, 1987). Additionally, some authors have established cut-offs of .20 in misfit plots (Kordoutis et al., 2021).

Since the model was retained, it was used in the remaining analyses instead of the 20-item version.

**Figure 1**

*Confirmatory factor analysis of the three-factor model of the IGD-10SV*



Note. F1 = Prevalence; F2 = Withdrawal; F3 = Craving.

**Table 5**  
*Correlation Residuals of the IGD-10SV*

<b>Indicator</b>	<b>6</b>	<b>13</b>	<b>20</b>	<b>4</b>	<b>10</b>	<b>16</b>	<b>11</b>	<b>12</b>	<b>14</b>	<b>15</b>
<b>Prevalence</b>										
Item 6	-									
Item 13	.07	-								
Item 20	<b>.13</b>	.06	-							
<b>Withdrawal</b>										
Item 4	.09	.01	.03	-						
Item 10	.08	.01	<b>.10</b>	.02	-					
Item 16	0	0	.02	.02	.01	-				
<b>Craving</b>										
Item 11	.05	0	.09	.04	.09	0	-			
Item 12	.05	<b>.10</b>	.08	.08	<b>.11</b>	.03	.01	-		
Item 14	0	.02	.02	.07	<b>.11</b>	.05	.04	.01	-	
Item 15	<b>.12</b>	.03	.04	.03	.09	.03	.06	.04	.08	-

## 2. Internal Consistency

The overall internal consistency of the IGD-10SV was good in this sample (see Table 6). Contrary to what Faraci et al. (2023) reported, all factors showed only modest internal consistency. Additionally, by excluding Item 13 in the factor Prevalence, both internal consistency coefficients would increase to .75. Consequently, since the value of internal consistency would remain almost the same, and excluding the item would lead to model underidentification (i.e., due to a lack of information in the sample it is impossible to estimate all of the free model parameters; Kline, 2023), the item was not excluded. The Internal Consistency of both CFA models and their factors is represented in Table 6 for comparison.

**Table 6**  
*Internal consistency of both scales and their factors*

<b>Tool</b>	<b>Factor</b>	<b><math>\alpha</math></b>	<b><math>\omega</math></b>
	Prevalence	.73	.74
IGD-10SV	Withdrawal	.72	.75

<b>Tool</b>	<b>Factor</b>	<b><math>\alpha</math></b>	<b><math>\omega</math></b>
	Craving	.62	.62
	Total	.83	.84
	Saliency	.62	.62
	Mood Modification	.74	.77
	Tolerance	.49	.61
IGD-20	Withdrawal Symptoms	.72	.75
	Conflict	.72	.74
	Relapse	.64	.71
	Total	.89	.90

*Note.* IGD-10SV = Internet Gaming Disorder-10 Short Version; IGD-20 = Internet Gaming Disorder Test

### 3. Convergent and Criterion Validity

The IGD-10SV showed good convergent validity as evidenced by its significant positive correlation with the GDT,  $r_{(377)} = .75, p < .001$ , and significant, but weak, positive correlations with both anxiety,  $r_{(377)} = .18, p < .001$ , and depression,  $r_{(377)} = .28, p < .001$ , scores of HADS.

Criterion validity was verified via the significant positive moderate correlation with daily gaming hours,  $r_{(377)} = .43, p < .001$ , and the significant positive, though weak correlation with weekly gaming days,  $r_{(377)} = .29, p < .001$ .

### 4. Test-retest Reliability

The scale showed a significant positive strong correlation with the retest application,  $r_{(110)} = .88, p < .001$ , supporting good test-retest reliability.

### 5. Hypotheses testing

The IGD-10SV total scores were significantly negatively correlated with the number of friends, however with a very weak relationship,  $\tau_b_{(377)} = -.08, p = .02$ .

Similarly, the IGD-10SV total scores were significantly negatively correlated with

the social gathering frequency, albeit with a very weak association,  $\tau_b(377) = -.11, p = .004$ .

Contrary to the remaining hypothesis, the total scores of the IGD-10SV were not negatively correlated with weekly hours of physical exercise,  $r(377) = -.04, p = .248$ .

## Discussion

This study aimed to validate the Internet Gaming Disorder Test (IGD-20; Pontes et al., 2014) for the Portuguese population as well as study the hypothesized negative associations between the scores of Internet Gaming Disorder (IGD) and the number of friends (NF), social gathering frequency (SGF) and weekly hours of physical exercise (PE). The results challenged the factorial structure of the IGD-20, in favor of a smaller, theoretically updated and culturally adapted scale. Notwithstanding that the IGD-10SV showed good overall internal consistency, convergent validity, criterion validity and test-retest reliability in a Portuguese sample, its factorial structure needs further scrutiny. Moreover, IGD presented very weak negative correlations with the NF and SGF but, contrary to the hypotheses, showed no connection with weekly PE.

In line with previous validation studies (e.g., L. I. Andrade et al., 2022; Cakiroglu & Soyulu, 2019; Faraci et al., 2023; Grajewski & Dragan, 2021; Hawi & Samaha, 2017; Yu et al., 2019), the Portuguese version of the IGD-20 struggled to fit within its original conceptual framework and showed very high latent factor intercorrelations. The results of the Confirmatory Factor Analysis (CFA) suggested that some of the concepts outlined initially for this scale tend to overlap and become almost undistinguishable. Moreover, the problems of local fit detected in the correlation matrix are believed to be caused by associations between items that (a) either do not have enough discriminant validity, or (b) are just naturally correlated. Consequently, some items from this scale should be revised to accommodate for a more parsimonious instrument. For instance, the salience and tolerance factors contain several items that pertain to “time” spent gaming, which could explain their systematic overlap.

Alternatively, IGD may not “fit” well enough within the conceptual framework it was initially based on. In other words, perhaps the latent factors of IGD behave differently compared to substance use addictions, leading to a six-factor model of addiction that does not correctly replicate. For example, Griffiths et al. (2016) have stated that:

The ‘tolerance’ criterion is clearly a consequence of modelling IGD criteria on that of substance disorder criteria and grounded in physiological reasons for requiring a greater intake. Consequently, this may not be as useful an indicator for problematic gaming as for other addictions. (p. 169)

The same authors also contend that other DSM-5 criteria (APA, 2013) are inadequate or poorly understood in their current formulation across several scales. Since the IGD-20 is heavily based on DSM-5 criteria, the mixed results from its subsequent validations might be indicative of those conceptual misunderstandings. One such example is item three: “I have significantly increased the amount of time I play games over last year.”. Belonging to the tolerance latent factor, item three showed the lowest pattern coefficient of the entire scale in this study. Therefore, in the context of IGD, tolerance might not manifest itself as increased time spent gaming, since at a certain point that is no longer possible (Griffiths et al., 2016; Ko, 2014). More item formulations and wording of the DSM-5 criteria (APA, 2013) are contested in several papers, as that is still an ongoing debate (Griffiths et al., 2016; Krossbakken et al., 2017; Kuss et al., 2017; Petry et al., 2014, 2016; Van Rooij & Kardefelt-Winther, 2017). Future studies could evaluate the items of the IGD-20 separately by combining exploratory factor analysis (EFA), CFA, and exploratory structural equation modelling (ESEM) techniques to rule out other potential problems with the scale. The ESEM allows for item intercorrelation, which is probably a more realistic premise than the more stringent CFA assumptions. Additionally, the works of Griffiths et al. (2016) and Van Rooij et al. (2017) discuss interesting suggestions that can aid in the challenging task of modifying or creating items.

Similarly to the Italian validation (Faraci et al., 2023), the internal consistency of the Portuguese IGD-10SV was good overall. However, the factor Craving showed only a modest internal consistency in this study. Moreover, that same factor was highly correlated with Withdrawal and shared correlation residuals with items 12 and 14. This could result from cultural or linguistic differences and should be properly investigated in future studies before research and especially clinical application in Portugal.

Additionally, just like previous studies have concluded, the IGD-10SV scores were positively correlated with anxiety and depression symptoms, even if showing weak positive correlations (Colder Carras et al., 2020; Ostinelli et al., 2021; Stavropoulos et al., 2022; Teng et al., 2021; Yu et al., 2019). The weaker correlations between IGD and symptoms of anxiety and depression could be the result of survey fatigue leading to less thought-out answers by the participants. Since the HADS was the last instrument presented

in a long questionnaire, both subscales could have suffered from this effect, leading to lower coefficients. The weaker correlations could also indicate discriminant validity. Unfortunately, without a clinical sample to compare the results to, it is not possible to assert whether anxiety and depression are more prevalent in disordered gamers than they are in the remaining population. Still, these results add to the convergent validity of the IGD-10SV and support the association between anxiety, depression and IGD. The very high association with the GDT also advocates for convergent validity of the scale. However, this assertion needs to be taken with caution as the GDT lacks a proper Portuguese validation. Their very high association in this sample shows that, at the very least, they measure analogous constructs. Other validation studies have reported similar high correlations between the IGD scale and measures of established problematic gaming or internet related behavior scales (Faraci et al., 2023; B.-N. Kim, 2019; Yu et al., 2019).

In order to further test this instrument's validity in Portugal, as well as the negative effects of IGD, future studies could implement additional measures of loneliness, stress, social interaction anxiety, self-esteem and perceived social support, as Faraci et al. (2023) have done.

The statistically significant positive and moderate correlation of the IGD-10SV with daily gaming hours is reminiscent of the criterion validity set by the Italian validation (Faraci et al., 2023). Together with the weaker, but statistically significant correlation with weekly gaming days, the results suggest a satisfactory criterion validity for this scale. However, as B.-N. Kim (2019) have theorized, gaming frequency may be too dependent on sample characteristics to be used as the best possible criterion for IGD. For example, gamers may play on specific days, such as weekends, for longer periods of time, while attending to other responsibilities during the rest of the week. This could explain the weaker correlation between IGD and weekly days engaged in gaming activities, as playing weekly requires juggling other psychosocial responsibilities for non-disordered gamers. Moreover, playing daily may not be related to any pathology whatsoever since gamers can play daily, but still not neglect any important aspects of their lives. Potentially, including a clinical sample of gamers with severe problematic gaming could reveal a higher association between both measures of gaming frequency, as the neglect of real-life responsibilities would be more prevalent. Alternatively, and as B.-N. Kim (2019) has outlined, "the core component of addiction might lie in the dysfunctional rather than excessive nature of engagement" (p. 5).

Additionally, if professional e-sports players or streamers are accounted for, more

days and longer time devoted to gaming activity would not necessarily be indicative of a higher correlation with problematic gaming. In that case, gaming would be considered a full-time job that requires a long time engaged in videogames, instead of a disorder (Maldonado-Murciano et al., 2022). This reasoning could also explain the findings, but since these gamers were not differentiated from the remaining participants in this study, it remains a conjecture. Until more stringent criteria and a “gold standard” is established in the area, this is the most widely defined approximation for criterion validity. Subsequently, it is considered that the IGD-10SV showed good enough criterion validity in this study.

Regarding test-retest reliability, the IGD-10SV showed good temporal stability for the eight days of difference between responses. Despite the good results, it is important to note that the interval between responses might have been too low, meaning that most participants probably still remembered their responses. Consequently, future researchers could try a 2-weeks interval between responses, like most similar instrument validations (King et al., 2020). This procedure supports that enough time has passed that the disordered behavior persists, but participants do not remember most of their responses. It is worth mentioning, however, that Marx et al. (2003) found no significant differences between the retest coefficients for 2-day and 2-weeks intervals.

Finally, our hypotheses were only partially verified in this sample. Both the measures of NF and SGF were negatively correlated with IGD scores, which seems to indicate that people with higher levels of problematic gaming have less friends and socialize less with their peers. These results are in line with Paulus et al.'s (2018) meta-analysis. Nevertheless, caution is warranted while interpreting these associations, as the effect sizes in the present study were very low. The lower correlations could be associated with the replacement of real-world friendships with similarly gratifying ones in the virtual world, like Sublette and Mullan (2012) have previously suggested. Additionally, both measures of friendship quality were probably lackluster in capturing other parts of the social world, such as satisfaction with friendships and social situations. While some people may consider having “many” friends and “going out” a lot as something positive, others might prefer lower amounts of both. Whether or not these relationships are fulfilling is probably more important in detecting the deleterious effects of IGD.

In this sample, the weekly hours of PE were not correlated with IGD scores, suggesting that the weekly practice of PE does not increase or decrease with the level of IGD. These results are in line with other studies that found no association between both variables (Bickham et al., 2013; Männikkö et al., 2015; Wack & Tantleff-Dunn, 2009).

One possible explanation for this phenomenon is that individuals without IGD are as sedentary as individuals with the disorder, in an ever-growing world where people are less physically active than recommended (OECD/WHO, 2023). In order to test this hypothesis, comparisons between clinical samples and control groups would be beneficial.

This study is not without limitations. Firstly, the data gathering protocol was lengthy, which could have had negative effects on the quality of the collected data. It is known that higher rates of non-responses, missing values or incomplete surveys increase with the length of the evaluation protocol (Chudoba, n.d.; Sharma, 2022). The HADS was probably the most affected instrument, because it was in the last section of the survey and no measures, such as section randomization, were taken to ensure response reliability. Future studies should consider smaller sections of the protocol to preserve the participants' interest and increase sample size. Researchers should also consider randomizing the instruments to control for participants' fatigue and order effects.

The forced-response format used in the survey may have also impacted the sample size (Stieger et al., 2007) as participants were forced to full completion and more people dropped. This decision was done with the intent of limiting the number of missing values; however, having a shorter survey would have been a better option.

The self-report nature of the protocol also presents biases well-known in the literature (e.g., social desirability, acquiescent responding, recall bias). These biases induce measurement error that may affect the reliability of the results (Kreitchmann et al., 2019). Future studies are advised to compare the results with those of clinical interviews and diagnostics to control for self-report bias. Alternatively, including social desirability scales may help in the identification of this type of bias.

In addition, the protocol was completely hosted online, which may generate self-selection bias and difficulty in describing the population (C. Andrade, 2020). The disadvantages of online surveys include the possibility for multiple responses, false demographic information, the requirement of internet connection and a device to participate. However, according to Wood and Griffiths (2007) these issues are alike those encountered in pen and paper surveys.

This study recruited a convenience sample and employed a non-probability snowball sampling method. Consequently, it is not possible to assert that it was representative or random, limiting the generalizability of the results. Moreover, despite the best efforts of the research team, gender, age and national representativeness were not achieved. Future studies should balance demographic characteristics to ensure a more

representative sample and generalizable results.

Furthermore, the CFA estimation method utilized for the three-factor model differed from the one used in the Italian validation. When measuring ordinal variables, while some argue that Maximum Likelihood (ML) or MLR perform well for five or more response levels (Rhemtulla et al., 2012), others advocate for the use of Weighted Least Squares Mean and Variance Adjusted (WLSMV), ULS or DWLS instead (Beauducel & Herzberg, 2006; Forero et al., 2009; Kořar & Yilmaz Kořar, 2015; Li, 2016), as it was done in this study. With the use of ML, the results might differ significantly. As such, additional analyses of these results or a thorough comparison and exploration of the factorial structure of the scale is important before further clinical or research application.

Due to time and planning constraints, this study could not deeply explore the factorial structure of the IGD-20 with more advanced and recent methods. Future researchers should first explore the scale via EFA and later confirm its results with another sample. Also, the global fit thresholds for model assessment were based on commonly established values, which are considered less reliable when the exact-fit test fails (Kline, 2023). Hence, an analysis with ESEM techniques, as previously stated, might be a good way to assess the factorial structure of the scale in the future. Additionally, including a clinical sample would be very important to establish a proper cut-off score for clinical screening and differentiate highly engaged gamers from disordered ones.

Thus, the current version is adapted the Portuguese population, but its validation is still an on-going process.

## **Conclusion**

Perhaps as a symptom of the conceptual ambiguities surrounding IGD, the validation studies of the IGD-20 seem to systematically struggle to adjust to its original conceptual model. In its current formulation, the items of the full scale do not seem to replicate sufficiently well across different cultures, suggesting that some dimensions and items are not well suited to measure the construct as they are currently presented.

In this study, the IGD-10SV seemed like an overall better option than the full 20-item scale when considering their fit to a factorial structure rooted in a theoretical framework of addiction. The smaller version of the scale showed moderate internal

consistency and good convergent validity, criterion validity, and test-retest reliability, supporting its possible adequacy in measuring IGD in the Portuguese population. Globally considered, the results of this study favour the revision of the items of the 20-item scale and its conceptual framework to establish a more adequate factorial structure.

Moreover, this study suggests that people with higher levels of IGD tend to report having less friends as well as meeting with them less often, but the IGD total scores were not associated with the practice of PE in this sample.

The IGD field of study needs a consensus over what aspects characterize the disorder and how to best assess them. One way of approaching a consensus, other than conceptual debate, is through the statistical evaluation of assessment scales and item performance across cultures. Despite its limitations, this study is a useful step towards that end.

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## **Appendices**

## Appendix A - District of Residence Frequencies

**Table 7**

*District of Residence frequencies*

District	Main Sample ( <i>n</i> = 556)		Subsample ( <i>n</i> = 379)		Retest Sample ( <i>n</i> = 112)	
	<i>n</i>	%	<i>n</i>	%	<i>n</i>	%
Porto	280	50.4	192	50.7	58	51.8
Lisbon	49	8.8	35	9.2	14	12.5
Faro	38	6.8	25	6.6	7	6.3
Viseu	31	5.6	23	6.1	7	6.3
Aveiro	28	5.0	19	5.0	2	1.8
Braga	25	4.5	19	5.0	10	8.9
Portalegre	20	3.6	11	2.9	2	1.8
Leiria	14	2.5	9	2.4	0	0
Setúbal	13	2.3	9	2.4	0	0
Madeira	11	2.0	9	2.4	2	1.8
Coimbra	8	1.4	2	0.5	0	0
Açores	6	1.1	4	1.1	2	1.8
Santarém	6	1.1	4	1.1	2	1.8
Viana do Castelo	6	1.1	3	0.8	1	0.9
Vila Real	6	1.1	4	1.1	1	0.9
Évora	4	0.7	3	0.8	1	0.9
Beja	3	0.5	2	0.5	1	0.9
Bragança	3	0.5	2	0.5	2	1.8
Guarda	3	0.5	3	0.8	0	0
Castelo Branco	2	0.4	1	0.3	0	0

## Appendix B - Full Protocol

- 1. Demographic questionnaire:** Age, Gender (M/F), District of Residence (dropdown list), Education Level (in years), Degree Enrolment, Enrolment Year, Academic Study Area, Faculty/School and University/Higher Education Institution, Number of Disapprovals and Civil Status.
- 2. Gaming related variables:** Gaming Platforms (i.e., Computer, Playstation, Wii, Xbox, Nintendo DS, PSP, Gameboy, Phone, Tablet, Others), Game Genres (i.e., Strategy, Simulation, Adventure, Platforming, Massive Multiplayer Online, Social Media, Luck-based, First-Person Shooter, Others), daily gaming hours, weekly gaming days, friend quantity, social gathering frequency and physical exercise weekly hours.
- 3. Gaming Addiction measures:** Internet Gaming Disorder Test (IGD-20; Pontes et al., 2014) and Gaming Disorder Test (GDT; Pontes et al., 2021).
- 4. Pathological Gambling measures:** South Oaks Gambling Screen (SOGS; Lesieur & Blume, 1987; European Portuguese adaptation by Lopes, 2008) and Gambling Disorder DSM-5 criteria (APA, 2013).
- 5. Psychopathy Measure:** Triarchic Psychopathy Measure (TriPM; Patrick et al., 2009; European Portuguese translation and adaptation by Vieira et al., 2014).
- 6. Anxiety and Depression Symptoms:** Hospital Anxiety and Depression Scale (Zigmond & Snaith, 1983; European Portuguese validation by Pais-Ribeiro et al., 2007)
- 7. Comorbidity:** Present and Past formally diagnosed psychopathology categories according to DSM-5 (APA, 2013).

## Appendix C - Ethics Committee Approval



ESCOLA  
SUPERIOR  
DE SAÚDE  
POLITÉCNICO  
DO PORTO

### PARECER DA COMISSÃO DE ÉTICA

Número de Registo da Comissão de Ética: **CE0100C** Data receção do Documento: 07/11/2022 Existência de entradas anteriores: Não

**TÍTULO DO TRABALHO:**

Validação Portuguesa do *Internet Gaming Disorder Test*: Relação com estilos de vida e quadros clínicos.

**INVESTIGADOR RESPONSÁVEL:**

Artemisa Rocha Dóres

**DATA PREVISTA PARA A REALIZAÇÃO DO TRABALHO:** Início 03/01/2023 Fim 31/12/2023

**RESUMO DO ESTUDO**

**OBJETIVOS:**

Nada a referir.

**AMOSTRA:**

Nada a referir.

**FORMULÁRIO DE DADOS A RECOLHER:**

Presentes e autorizado a sua utilização pelos seus próprios autores. Nada a referir.

**MATERIAL:**

Nada a referir.

**MÉTODOS:**

Nada a referir.

**RISCOS:**

Nada a referir.

**CONSENTIMENTO INFORMADO:**

Solicitado no início do questionário

**AUTORIZAÇÃO PELOS RESPONSÁVEIS LOCAIS:**

Presentes. Nada a referir.

**APRECIÇÃO DA COMISSÃO DE ÉTICA:**

Processo bem instruído. Apresenta o link para consulta e análise do questionário.

**PARECER FINAL DA COMISSÃO DE ÉTICA:**

De acordo com todos os dados analisados, o parecer é "favorável", desde que cumpridas todas as diretrizes submetidas a esta Comissão, recomendando-se que a decisão seja suspensa caso haja algum incumprimento grave.

Assinado por: **PEDRO MANUEL RIBEIRO DA  
ROCHA MONTEIRO**  
Num. de Identificação: 09132856  
Data: 2022.12.07 13:25:04 +0000

06/12/2022



SGS ESS.004.M0.318.02

## Appendix D - Correlation Comparisons

**Table 8**  
*Correlation Comparisons*

Variable	IGD-10SV					
	Outliers			No Outliers		
	<i>r</i>	<i>r<sub>s</sub></i>	<i>df</i>	<i>r</i>	<i>r<sub>s</sub></i>	<i>df</i>
GDT	.75***	.71***	377	.71***	.70***	366
HADS_A	.18***	.18***	377	.17***	.18***	376
HADS_D	.28***	.30***	377	.29***	.30***	372
Daily gaming hours	.43***	.50***	377	.48***	.50***	368
Weekly gaming days	.29***	.30***	377	.29***	.30***	376
IGD-10SV_II	.88***	.87***	110			
Weekly hours of physical exercise	-.04	-.04	377	-.04	-.04	367

*Note.* GDT = Gaming Disorder Test; HADS\_A = anxiety subscale of HADS; HADS\_D = depression subscale of HADS; IGD-10SV = Internet Gaming Disorder-10 Short Version, IGD-10SV\_II = Retest phase of the IGD-10SV.

\*\*\* $p < .001$

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